

Lecture 2b: Survey weights

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Advanced Methods of Social Research (SOCL 420)

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Source: Treiman, Donald J. 2009. Quantitative Data Analysis: Doing Social Research to Test Ideas. San Francisco: Jossey-Bass. Chapter 9 (pp. 195–224).



Outline

- Inferential statistics
- Survey weights
- Weight options in Stata
- Complex sample cluster design
- Weights in the General Social Survey (GSS)
- Examples of descriptive statistics



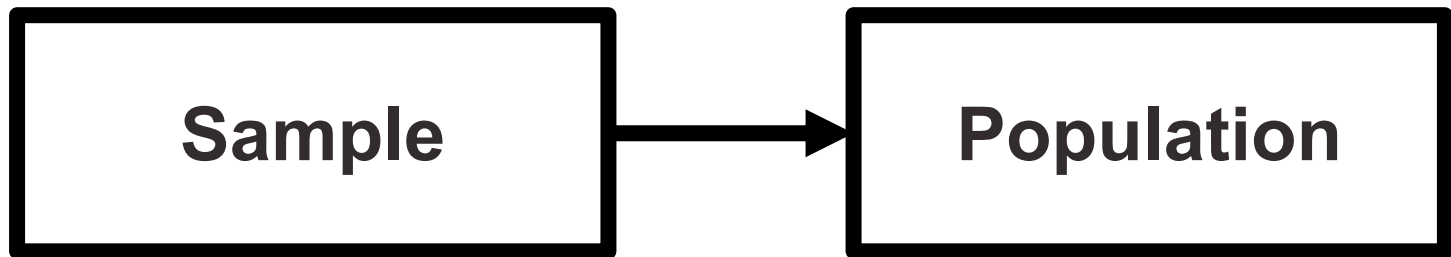
Inferential statistics

- Social scientists need inferential statistics
 - They almost never have the resources or time to collect data from every case in a population
- Inferential statistics uses data from samples to make generalizations about populations
 - **Population** is the total collection of all cases in which the researcher is interested
 - **Samples** are carefully chosen subsets of the population
- With proper techniques, generalizations based on samples can represent populations

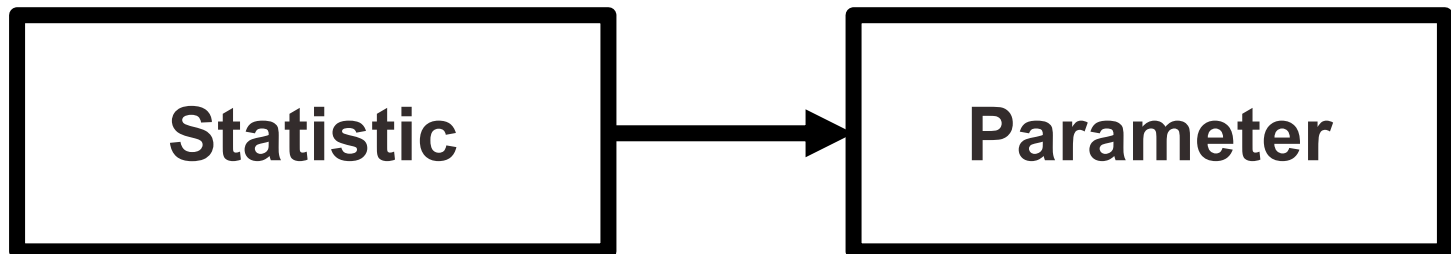


Basic logic and terminology

- Information from samples is used to estimate information about the population



- Statistics: characteristics of samples
- Parameters: characteristics of populations
- Statistics are used to estimate parameters





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Survey weights

| Name | Number of observations collected in the survey | Weight to expand to population size | Weight to maintain sample size |
|-------|--|-------------------------------------|--------------------------------|
| José | 1 | 4 | 0.8 |
| Maria | 1 | 6 | 1.2 |
| Total | 2 | 10 | 2 |

Survey weight =

Population weight * (Sum of survey weights / Sum of population weights)



Weights for tables

- When we use a sample to estimate the absolute number of people
 - For an area
 - For a specific sub-group
 - We use weights to expand to population size
- If we use a sample to estimate the proportion of people in a specific sub-group
 - And we are not concerned with the absolute value
 - We use weights to maintain the sample size (we focus on percentages)

Weights for regressions

- In a simple linear regression, the test of statistical significance for a β coefficient (t -test) is estimated as

$$t = \frac{\hat{\beta}}{SE_{\hat{\beta}}} = \frac{\hat{\beta}}{\sqrt{\frac{MSE}{S_{xx}}}} = \frac{\hat{\beta}}{\sqrt{\frac{RSS}{df * S_{xx}}}} = \frac{\hat{\beta}}{\sqrt{\frac{\sum_i (y_i - \hat{y}_i)^2}{(n - 2) \sum_i (x_i - \bar{x})^2}}}$$

- $SE_{\hat{\beta}}$: standard error of β
- MSE : mean squared error = RSS / df
- RSS : residual sum of squares = $\sum_i (y_i - \hat{y}_i)^2 = \sum_i \hat{e}_i^2$
- df : degrees of freedom = $n-2$ for simple linear regression
 - 2 statistics (slope and intercept) are estimated to calculate sum of squares
- S_{xx} : corrected sum of squares for x (total sum of squares)



Weights for regressions

- If we use a weight that expands to the population size (N) on regressions
 - We would be incorrectly informing the statistical software that we have a sample with enormous size
 - This would artificially increase the test of statistical significance for the coefficient

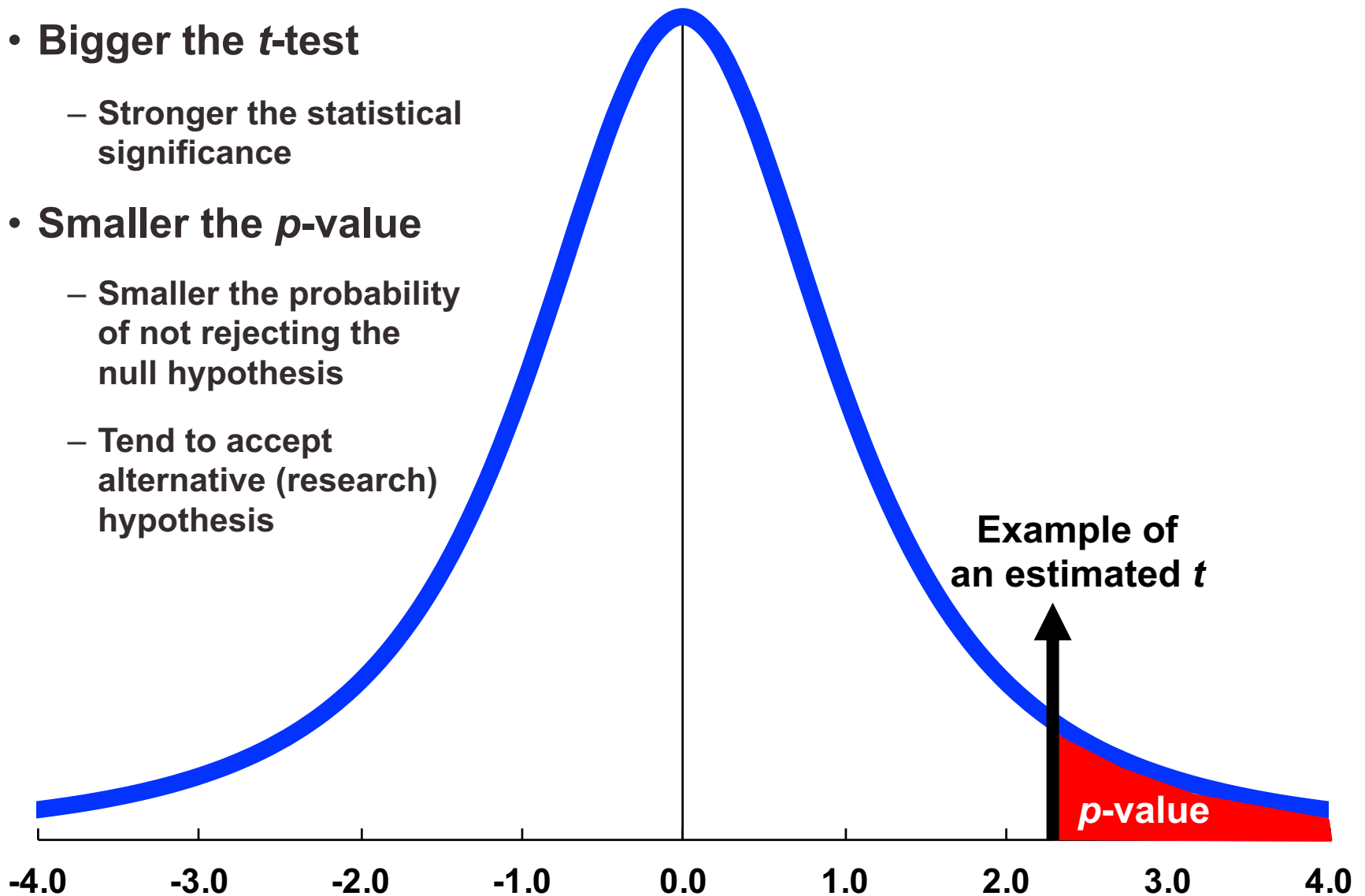
$$\uparrow t = \frac{\hat{\beta}}{SE_{\hat{\beta}}} = \frac{\hat{\beta}}{\sqrt{\frac{MSE}{S_{xx}}}} = \frac{\hat{\beta}}{\sqrt{\frac{MSE}{S_{xx}}}} = \frac{\hat{\beta}}{\sqrt{\frac{\sum_i (y_i - \hat{y}_i)^2}{(n-2) \sum_i (x_i - \bar{x})^2}}}$$

The equation shows the derivation of the t-statistic for a regression coefficient. A red arrow points up to the t-statistic, and another red arrow points down to the denominator of the final fraction.

- We have to inform the weight related to the sample design, but we should maintain the sample size (n)

t distribution ($df = 2$)

- **Bigger the t -test**
 - Stronger the statistical significance
- **Smaller the p -value**
 - Smaller the probability of not rejecting the null hypothesis
 - Tend to accept alternative (research) hypothesis



Decisions about hypotheses

| Hypotheses | $p < \alpha$ | $p > \alpha$ |
|-------------------------------------|--------------|---------------|
| Null hypothesis (H_0) | Reject | Do not reject |
| Alternative hypothesis (H_1) | Accept | Do not accept |

– ***p*-value** is the probability of not rejecting the null hypothesis

– If a statistical software gives only the two-tailed *p*-value, divide it by 2 to obtain the one-tailed *p*-value

| Significance level (α) | Confidence level (success rate) |
|------------------------------------|------------------------------------|
| 0.10 (10%) | 90% |
| 0.05 (5%) | 95% |
| 0.01 (1%) | 99% |
| 0.001 (0.1%) | 99.9% |





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Weight options in Stata

- Frequency weight (fweight)
- "Importance" weight (iweight)
- Analytic weight (aweight)
- Sampling weight (pweight)



Extract of 2018 ACS microdata

| | year | strata | cluster | perwt | hhwt | sex | age | income |
|----|------|--------|--------------|--------|--------|--------|-----|--------|
| 1 | 2018 | 360248 | 2.018012e+12 | 56.00 | 56.00 | Male | 46 | 28000 |
| 2 | 2018 | 360248 | 2.018012e+12 | 51.00 | 51.00 | Male | 20 | 5000 |
| 3 | 2018 | 360248 | 2.018012e+12 | 76.00 | 76.00 | Female | 84 | 0 |
| 4 | 2018 | 360248 | 2.018012e+12 | 55.00 | 55.00 | Female | 18 | 1200 |
| 5 | 2018 | 360248 | 2.018012e+12 | 143.00 | 143.00 | Female | 56 | 1500 |
| 6 | 2018 | 360248 | 2.018012e+12 | 198.00 | 198.00 | Male | 31 | 10000 |
| 7 | 2018 | 360248 | 2.018012e+12 | 48.00 | 48.00 | Female | 19 | 2000 |
| 8 | 2018 | 360248 | 2.018012e+12 | 48.00 | 48.00 | Male | 25 | 7000 |
| 9 | 2018 | 360248 | 2.018012e+12 | 65.00 | 65.00 | Female | 18 | 0 |
| 10 | 2018 | 360248 | 2.018012e+12 | 53.00 | 53.00 | Female | 18 | 15000 |
| 11 | 2018 | 360248 | 2.018012e+12 | 17.00 | 17.00 | Male | 63 | 0 |
| 12 | 2018 | 360248 | 2.018012e+12 | 39.00 | 39.00 | Female | 18 | 4000 |
| 13 | 2018 | 360248 | 2.018012e+12 | 104.00 | 104.00 | Male | 21 | 1000 |
| 14 | 2018 | 360248 | 2.018012e+12 | 200.00 | 200.00 | Male | 40 | 80000 |
| 15 | 2018 | 360248 | 2.018012e+12 | 20.00 | 20.00 | Male | 33 | 0 |
| 16 | 2018 | 360248 | 2.018012e+12 | 59.00 | 59.00 | Male | 19 | 2900 |
| 17 | 2018 | 360248 | 2.018012e+12 | 56.00 | 56.00 | Male | 55 | 0 |
| 18 | 2018 | 360248 | 2.018012e+12 | 77.00 | 77.00 | Male | 18 | 9000 |
| 19 | 2018 | 360248 | 2.018012e+12 | 16.00 | 16.00 | Female | 41 | 1100 |
| 20 | 2018 | 360248 | 2.018012e+12 | 46.00 | 46.00 | Male | 33 | 0 |

Frequency weight

- **FWEIGHT**

- Expands survey size to the population size
- Indicates the number of duplicated observations
- Used on tables to generate frequencies
- Can be used in frequency distributions only when weight variable is discrete (no fractional numbers)

```
tab x [fweight = weight]
```



"Importance" weight

- **IWEIGHT**

- Indicates the "importance" of the observation in some vague sense
- Has no formal statistical definition
- Any command that supports iweights will define exactly how they are treated
- Intended for use by programmers who want to produce a certain computation
- Can be used in frequency distributions even when weight variable is continuous (fractional numbers)

```
tab x [iweight = weight]
```



Analytic weight

- **AWEIGHT**

- Inversely proportional to the variance of an observation
- Variance of the j th observation is assumed to be σ^2/w_j , where w_j are the weights
- For most Stata commands, the recorded scale of aweights is irrelevant
- Stata internally rescales frequencies, so sum of weights equals sample size

```
tab x [aweight = weight]
```

```
regress y x1 x2 [aweight = weight]
```



More about analytic weight

- Observations represent averages and weights are the number of elements that gave rise to the average

| group | x | y | n |
|-------|-----|------|---|
| 1 | 3.5 | 26.0 | 2 |
| 2 | 5.0 | 20.0 | 3 |

- Instead of

| group | x | y |
|-------|---|----|
| 1 | 3 | 22 |
| 1 | 4 | 30 |
| 2 | 8 | 25 |
| 2 | 2 | 19 |
| 2 | 5 | 16 |

- Usually, survey data is collected from individuals and households (not as averages)
 - Thus, aweights are not appropriate for most cases



Sampling weight

- **PWEIGHT**

- Denote the inverse of the probability that the observation is included due to the sampling design
- Variances, standard errors, and confidence intervals are estimated with a more precise procedure
- Indicated for statistical regressions to estimate robust standard errors
 - Obtain unbiased standard errors of OLS coefficients under heteroscedasticity (i.e., residuals not randomly distributed)
 - Robust standard errors are usually larger than conventional ones

```
regress y x1 x2 [pweight = weight]
```



Summary of Stata weights

WEIGHTS IN FREQUENCY DISTRIBUTIONS

| Weight unit of measurement | Expand to population size | Maintain sample size |
|----------------------------|---------------------------|----------------------|
| Discrete | fweight | aweight |
| Continuous | iweight | |

WEIGHTS IN STATISTICAL REGRESSIONS should maintain sample size

| Robust standard error | Adjusted R ² , TSS, ESS, RSS |
|---|--|
| pweight reg y x, vce(robust) reg y x, vce(cluster area) | aweight outreg2 |



Example of 2018 ACS weight

. sum perwt, d

Person weight

| Percentiles | | Smallest | | |
|-------------|------------|-------------|-------------|------------------|
| 1% | 10 | 1 | | |
| 5% | 19 | 1 | | |
| 10% | 29 | 1 | Obs | 3,214,539 |
| 25% | 52 | 1 | Sum of wgt. | 3,214,539 |
| 50% | 80 | | Mean | 101.7774 |
| | | | Std. dev. | 83.93534 |
| | | Largest | | |
| 75% | 124 | 1916 | | |
| 90% | 195 | 1990 | Variance | 7045.14 |
| 95% | 263 | 2097 | Skewness | 2.845116 |
| 99% | 427 | 2313 | Kurtosis | 17.99265 |

Example of 2018 ACS weight

. tab sex

| Sex | Freq. | Percent | Cum. |
|--------|-----------|---------|--------|
| Male | 1,574,618 | 48.98 | 48.98 |
| Female | 1,639,921 | 51.02 | 100.00 |
| Total | 3,214,539 | 100.00 | |

. tab sex [fweight=perwt]

| Sex | Freq. | Percent | Cum. |
|--------|-------------|---------|--------|
| Male | 161,072,404 | 49.23 | 49.23 |
| Female | 166,095,035 | 50.77 | 100.00 |
| Total | 327,167,439 | 100.00 | |

. tab sex [iweight=perwt]

| Sex | Freq. | Percent | Cum. |
|--------|-------------|---------|--------|
| Male | 161,072,404 | 49.23 | 49.23 |
| Female | 166,095,035 | 50.77 | 100.00 |
| Total | 327,167,439 | 100.00 | |

. tab sex [aweight=perwt]

| Sex | Freq. | Percent | Cum. |
|--------|-----------|---------|--------|
| Male | 1,582,595 | 49.23 | 49.23 |
| Female | 1,631,944 | 50.77 | 100.00 |
| Total | 3,214,539 | 100.00 | |



Example of 2021 GSS weight

. sum wtssnrps, d

person post-stratification weight, nonrespondents
adjusted

| | Percentiles | Smallest | | |
|-----|-------------|----------|-------------|----------|
| 1% | .243687 | .1723802 | | |
| 5% | .30024 | .1738938 | | |
| 10% | .4057674 | .1926333 | Obs | 4,032 |
| 25% | .5423563 | .2104285 | Sum of wgt. | 4,032 |
| 50% | .8183308 | | Mean | 1 |
| | | Largest | Std. dev. | .7260472 |
| 75% | 1.212269 | 6.51434 | | |
| 90% | 1.798724 | 6.903664 | Variance | .5271445 |
| 95% | 2.27083 | 7.218392 | Skewness | 2.825826 |
| 99% | 3.986099 | 7.557038 | Kurtosis | 15.89999 |

Example of 2021 GSS weight

```
. tab sex, m
```

| respondents sex | Freq. | Percent | Cum. |
|--------------------|-------|---------|--------|
| male | 1,736 | 43.06 | 43.06 |
| female | 2,204 | 54.66 | 97.72 |
| .i | 19 | 0.47 | 98.19 |
| .n | 71 | 1.76 | 99.95 |
| .s | 2 | 0.05 | 100.00 |
| Total | 4,032 | 100.00 | |

```
. tab sex [fweight=wtssnrps], m
may not use noninteger frequency weights
r(401);
```

```
. tab sex [iweight=wtssnrps], m
```

| respondents sex | Freq. | Percent | Cum. |
|--------------------|-------------|---------|--------|
| male | 1,904.2566 | 47.23 | 47.23 |
| female | 1,993.21543 | 49.43 | 96.66 |
| .i | 18.1122752 | 0.45 | 97.11 |
| .n | 113.299832 | 2.81 | 99.92 |
| .s | 3.11586052 | 0.08 | 100.00 |
| Total | 4,032 | 100.00 | |

```
. tab sex [aweight=wtssnrps], m
```

| respondents sex | Freq. | Percent | Cum. |
|--------------------|-------------|---------|--------|
| male | 1,904.2566 | 47.23 | 47.23 |
| female | 1,993.21543 | 49.43 | 96.66 |
| .i | 18.1122752 | 0.45 | 97.11 |
| .n | 113.299832 | 2.81 | 99.92 |
| .s | 3.11586052 | 0.08 | 100.00 |
| Total | 4,032 | 100.00 | |



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Complex sample cluster design

- To calculate standard errors correctly, variables for sample cluster design must be used
 - Without design variables, Stata will assume a simple random sample and underestimate standard errors
- Strata are created based on the lowest level of geography available in each sample
 - We use additional statistical techniques that account for the complex sample design to produce correct standard errors and statistical tests



Cluster design for tables

- If we want to estimate a confidence interval for a sample statistic (mean or proportion), we need to inform the complex survey design
- **Confidence interval** is a range of values used to estimate the true population parameter
- **Confidence level** is the success rate of the procedure to estimate the confidence interval
- Larger confidence levels generate larger confidence intervals



Confidence level, α , and Z

| Confidence level (1 - α) * 100 | Significance level alpha (α) | $\alpha / 2$ | Z score |
|---|--|--------------|------------------------------|
| 90% | 0.10 | 0.05 | ± 1.65 |
| 95% | 0.05 | 0.025 | ± 1.96 |
| 99% | 0.01 | 0.005 | ± 2.58 |
| 99.9% | 0.001 | 0.0005 | ± 3.32 |
| 99.99% | 0.0001 | 0.00005 | ± 3.90 |



Confidence intervals from samples

c.i. = sample estimate \pm margin of error

*c.i. = sample estimate \pm score of confidence level * standard error*

- Sample mean (\bar{x}), standard deviation (s), $n < 30$

$$c.i. = \bar{x} \pm t \left(\frac{s}{\sqrt{n}} \right) \quad df = n - 1$$

- Sample mean (\bar{x}), standard deviation (s), $n \geq 30$

$$c.i. = \bar{x} \pm Z \left(\frac{s}{\sqrt{n - 1}} \right)$$

- Sam. proportion (P_s), pop. proportion (P_u), $n \geq 30$

$$c.i. = P_s \pm Z \sqrt{\frac{P_u(1 - P_u)}{n}}$$



Cluster design for regressions

- We also need to inform cluster design for regressions, because the t -test utilizes standard errors

$$t = \frac{\hat{\beta}}{SE_{\hat{\beta}}} = \frac{\hat{\beta}}{\sqrt{\frac{MSE}{S_{xx}}}} = \frac{\hat{\beta}}{\sqrt{\frac{RSS}{df * S_{xx}}}} = \frac{\hat{\beta}}{\sqrt{\frac{\sum_i (y_i - \hat{y}_i)^2}{(n - 2) \sum_i (x_i - \bar{x})^2}}}$$

- $SE_{\hat{\beta}}$: standard error of $\hat{\beta}$
- MSE : mean squared error = RSS / df
- RSS : residual sum of squares = $\sum_i (y_i - \hat{y}_i)^2 = \sum_i \hat{e}_i^2$
- df : degrees of freedom = $n - 2$ for simple linear regression
- S_{xx} : corrected sum of squares for x (total sum of squares)

Cluster design & standard error

- Sample cluster designs underestimate standard errors, because they tend to select individuals with more similar characteristics from the same clusters
 - Simple random samples would provide more variation (higher standard errors), because they give the same chance of selection for all individuals in the population
- When we inform the cluster design, the standard error tends to increase and statistical significance decreases

$$\downarrow t = \frac{\hat{\beta}}{\uparrow SE_{\hat{\beta}}} = \frac{\hat{\beta}}{\sqrt{\frac{MSE}{S_{xx}}}} = \frac{\hat{\beta}}{\sqrt{\frac{RSS}{df * S_{xx}}}} = \frac{\hat{\beta}}{\sqrt{\frac{\sum_i (y_i - \hat{y}_i)^2}{(n - 2) \sum_i (x_i - \bar{x})^2}}}$$



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Weights in GSS

- The General Social Survey (GSS) targets the adult population (18+) living in U.S. households
- Due to the adoption of the sub-sampling design of non-respondents, a weight must be employed when using the GSS 2004 and after
- There are three continuous weight variables
 - WTSS
 - WTSSNR
 - WTSSALL
- They all maintain the original sample size, even in frequency distributions with "iweight"



WTSS

- WTSS variable takes into consideration
 - Sub-sampling of non-respondents
 - Number of adults in the household
- In years prior to 2004, a value of one is assigned to all cases, so they are effectively unweighted
 - Number of adults can be utilized to make this adjustment for years prior to 2004



WTSSNR

- WTSSNR variable takes into consideration
 - Sub-sampling of non-respondents
 - Number of adults in the household
 - Differential non-response across areas
- In years prior to 2004, a value of one is assigned to all cases, so they are effectively unweighted
 - Number of adults can be utilized to make this adjustment for years prior to 2004
 - Area non-response adjustment is not possible



WTSSALL

- WTSSALL takes WTSS and applies an adult weight to years before 2004
- The weight value of WTSSALL is the same as WTSS for 2004 and after
- Researchers who use the GSS data before or after 2004 may consider using the WTSSALL weight variable

```
tab x [aweight = wtssall]
```

```
sum x [aweight = wtssall]
```



GSS has a cluster sample

(https://gssdataexplorer.norc.umd.edu/gss_stdError)

- First- and second-stage units are selected with probabilities proportional to size
 - Size is defined by number of housing units
- Third-stage units (housing units) are selected to be an equal-probability sample
 - This results in roughly the same number of housing units selected per second-stage sampling unit



GSS variables for cluster design

(https://gssdataexplorer.norc.umd.edu/gss_stdError)

- There are two design variables
 - VSTRAT
 - VPSU
- First-stage unit
 - VSTRAT: Variance Stratum
 - National Frame Areas (NFAs): one or more counties
- Second-stage unit
 - VPSU: Variance Primary Sampling Unit
 - Segments: block, group of blocks, or census tract



GSS complex sample design

(https://gssdataexplorer.norc.org/gss_stdError)

- Account for GSS sample design in Stata

```
svyset [weight=wtssall], strata(vstrat) psu(vpsu) singleunit(scaled)
```

- After "svyset," you should indicate survey design with the option "svy" for commands that estimate standard errors

```
svy: mean y
```

```
svy: reg y x1 x2
```



Strata with single sampling unit

(https://gssdataexplorer.norc.org/gss_stdError)

- VSTRAT and VPSU were created with a minimum of three respondents within a cell
 - If all cases are missing on a variable, you get an error message in Stata
 - "Missing standard error because of stratum with single sampling unit"
- It is recommended to utilize the "subpop" option for any subdomain analyses (e.g., for males)

```
svy, subpop(if sex==1): tab x
```

- You can also specify that strata with one sampling unit are "centered" at grand mean instead of stratum mean

```
svyset [weight=wtssall], strata(vstrat) psu(vpsu) singleunit(centered)
```





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Example: 2021 GSS in Stata (nominal-level variable)

```
. tab sex
```

| respondents sex | Freq. | Percent | Cum. |
|--------------------|-------|---------|--------|
| male | 1,736 | 44.06 | 44.06 |
| female | 2,204 | 55.94 | 100.00 |
| Total | 3,940 | 100.00 | |

```
. tab sex [iweight=wtssnrps]
```

| respondents sex | Freq. | Percent | Cum. |
|--------------------|-------------|---------|--------|
| male | 1,904.2566 | 48.86 | 48.86 |
| female | 1,993.21543 | 51.14 | 100.00 |
| Total | 3,897.472 | 100.00 | |

```
. svyset [weight=wtssnrps], strata(vstrat) psu(vpsu) singleunit(scaled)
(sampling weights assumed)
```

```
. svy: tab sex
(running tabulate on estimation sample)
```

```
Number of strata = 9
Number of PSUs = 3,492
Number of obs = 3,940
Population size = 3,897.472
Design df = 3,483
```

| responden ts sex | proportion |
|---------------------|------------|
| male | .4886 |
| female | .5114 |
| Total | 1 |

Key: proportion = Cell proportion

Example: 2021 GSS in Stata (ordinal-level variable)

`. tab degree`

| r's highest degree | Freq. | Percent | Cum. |
|--------------------------|-------|---------|--------|
| less than high school | 246 | 6.14 | 6.14 |
| high school | 1,597 | 39.84 | 45.97 |
| associate/junior college | 370 | 9.23 | 55.20 |
| bachelor's | 1,036 | 25.84 | 81.04 |
| graduate | 760 | 18.96 | 100.00 |
| Total | 4,009 | 100.00 | |

`. tab degree [iweight=wtssnrps]`

| r's highest degree | Freq. | Percent | Cum. |
|--------------------------|-------------|---------|--------|
| less than high school | 480.972702 | 11.99 | 11.99 |
| high school | 1,891.6334 | 47.15 | 59.13 |
| associate/junior college | 452.656901 | 11.28 | 70.42 |
| bachelor's | 681.8664156 | 16.99 | 87.41 |
| graduate | 505.084448 | 12.59 | 100.00 |
| Total | 4,012.2139 | 100.00 | |

```
. svyset [weight=wtssnrps], strata(vstrat) psu(vpsu) singleunit(scaled)
(sampling weights assumed)
```

`. svy: tab degree`

(running **tabulate** on estimation sample)

```
Number of strata = 9
Number of PSUs = 3,543
Number of obs = 4,009
Population size = 4,012.2139
Design df = 3,534
```

| r's highest degree | proportion |
|--------------------------|------------|
| less than high school | .1199 |
| high school | .4715 |
| associate/junior college | .1128 |
| bachelor's | .1699 |
| graduate | .1259 |
| Total | 1 |

Key: proportion = Cell proportion

Example: 2021 GSS in Stata (interval-ratio-level variable)

```
. sum conrinc
```

| Variable | Obs | Mean | Std. dev. | Min | Max |
|----------|-------|----------|-----------|-----|----------|
| conrinc | 2,456 | 41722.79 | 39243.69 | 336 | 170912.6 |

```
. sum conrinc [iweight=wtssnrps]
```

| Variable | Obs | Weight | Mean | Std. dev. | Min | Max |
|----------|-------|------------|----------|-----------|-----|----------|
| conrinc | 2,456 | 2453.15509 | 37647.74 | 37376.88 | 336 | 170912.6 |

```
. svy: mean conrinc
```

(running mean on estimation sample)

Survey: Mean estimation

```
Number of strata = 9          Number of obs = 2,456
Number of PSUs   = 2,241     Population size = 2,453.1551
Design df        =          Design df = 2,232
```

| | Mean | Linearized std. err. | [95% conf. interval] | |
|---------|----------|-------------------------|----------------------|----------|
| conrinc | 37647.74 | 850.3902 | 35980.1 | 39315.38 |

```
. estat sd
```

| | Mean | Std. dev. |
|---------|----------|-----------|
| conrinc | 37647.74 | 37376.87 |



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